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Shivaji Shripati Desai Department of Statistics, Gopal Krishna Gokhale College, Kolhapur (MS), India., ssd.stats@gmail.com

D N. Kashid Shivaji University, Kolhapur, Maharashtra, India., dnk\_stats@unishivaji.ac.in

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# **Cover Page Footnote**

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# Support Vector Machine-based Modified Sp Statistic for Subset Selection with Non-Normal Error Terms

**Shivaji Shripati Desai** Gopal Krishna Gokhale College, Kolhapur, India **D N. Kashid** Shivaji University, Kolhapur, India

Support vector machine (SVM) is used for estimation of regression parameters to modify the sum of cross products (Sp). It works well for some nonnormal error distributions. The performance of existing robust methods and the modified Sp is evaluated through simulated and real data. The results show the performance of the modified Sp is good.

*Keywords*: support vector machine (SVM), support vector regression (SVR), robust subset selection, Sp- statistic, SSp- statistic

# Introduction

Regression models are used in almost all fields for establishing the relationship between a variable and a set of variables. Such models are useful tools to predict future values of the response variable given the values of predictors. The multiple linear regression model is given by

$$Y = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{e} \tag{1}$$

where Y is  $n \times 1$  vector of observations on response variable, X is a known  $n \times k$  matrix of observations on  $k \times 1$  predictors with 1's in the first column,  $\hat{a} = (\beta_0, \beta_1, \beta_2, \dots, \beta_{k-1})$ ' is  $k \times 1$  vector of unknown regression parameters and e is  $n \times 1$  vector of unknown errors. The assumptions on the model in (1) are E(e)=0,  $Cov(e) = \sigma^2 I Cov(e) = \sigma^2 I$  and  $e \sim N_n(0, \sigma^2 I)$  where, I is the identity matrix of order  $n \times n$  and  $\sigma^2$  is error variance.

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Important steps in regression are to obtain the estimates of regression parameters, choose an appropriate model for the given data and to predict the future values of response variable as accurate as possible. Least squares (LS) method is generally used for estimation of parameters in linear regression. The performance of least squares method is excellent if underlying assumptions are true, while it deteriorates when the data contains an outlier and (or) the distribution of error variable is non-normal. Huber's (1981) M-estimator is usually used when error distribution is nonnormal but it is close to normal, and Jaeckel's (1972) rank-based estimator performs well for almost any possible distribution of error (Birkes & Dodge, 1993, pp. 111).

In the set of possible predictors to be included in the model, some of them may be redundant and are required to be eliminated based on the observed data, which will give sufficient predictive accuracy. This is popularly known as subset selection in regression. Miller (2002) indicated fitting a model with a large number of predictors is neither economical nor practicable and in practice usually a model based on a small subset of predictors gives more accurate predictions.

There is considerable literature on subset selection methods in regression (Hocking, 1976; Thompson, 1978a, b; Rao & Wu, 1989). Standard texts such as Miller (2002), Draper and Smith (2003) and Montgomery et al. (2006) provided a good description of subset selection methodologies.

The majority of subset selection methods, including the Cross product (Cp) criterion (Mallow, 1973), are based on the LS estimator of  $\beta$ . In view of the performance of LS estimator in the presence of outlier and nonnormal error distribution, the subset selection methods based on such estimates will select the wrong subset, as demonstrated in Ronchetti and Staudte (1994).

Many methods were proposed for the choice of a subset by minimizing a criterion, including those by Akaike (1973), Schwartz (1978), Shibata (1984), Rao and Wu (1989), and Kundu and Murali (1996). Some robust subset selection procedures were proposed, including the Robust AIC (Ronchetti, 1985) and Robust versions of Mallow's (1973) Cp, called RCp (Ronchetti & Staudte, 1994). Kashid and Kulkarni (2002) suggested a more general Sum of Cross products (Sp) criterion for subset selection. It uses scaled difference between robust predicted values from subset and full model to perform subset selection in linear regression. Their results showed performance of Sp is better than Cp in the presence of outliers. Baierl et al. (2007) proposed a robust version of BIC based on Huber's M-estimator and called it as Robust BIC.

These M-estimators works well under certain assumptions. They are not guaranteed to produce a good subset selection if the data do not support the underlying assumptions. An alternative is to base a subset selection procedure on a data dependent prediction method such as the Support Vector Machine (SVM), the focus of this study. It is a growing area in machine learning introduced by Boser et al. (1992) in COLT. The basic task of SVM is to explore data (input-output pairs) and provide optimally accurate predictions for unseen data (Nalbantov, 2003). Vapnik et al. (1997) used SVM for regression and called it as Support Vector Regression (SVR). The SVR problem is formulated as a convex optimization problem which has the advantage of being free from local minima. SVMs based on the  $\varepsilon$  insensitive loss for regression are consistent and robust even for heavy-tailed distributions (Christmann et al., 2008, 2009; Messem & Christmann, 2010). SVR is used here for parameter estimation and to obtain predicted values from the full model and subset models.

# Support Vector Regression

In SVR, an unknown regression function  $f(x_i)$  based on data set  $(x_i, y_i)$ , i=1,2,...,n of input vectors  $x_i \in \mathbb{R}^{k-1}$  (i<sup>th</sup> row of design matrix X excluding first element 1) and associated target  $y_i \in \mathbb{R}$ , is estimated in the form,

$$y_i = f(\mathbf{x}_i) + e_i \tag{2}$$

where  $e_i$  is error term.

For linear regression, the function can be written as,

$$f(\mathbf{x}_i) = b + \mathbf{x}_i \mathbf{w} \,, \tag{3}$$

where  $w = (w_1, w_2, \dots, w_{k-1}) \in \mathbb{R}^{k-1}, b \in \mathbb{R}$  is a bias and  $x_i w$  is a dot product of  $x_i$  and w.

Thus, Equation (2) becomes,

$$y_i = b + x_i w + e_i$$
,  $i = 1, 2, \dots, n$ .

In matrix notations,

$$Y = X\beta + e_{,}$$

where  $\hat{a} = (b, w_1, w_2, \dots, w_{k-1})'$ , Y, X and e are the same as defined in Equation (1). The above equation is equivalent to Equation (1).

Using the  $\varepsilon$  insensitive loss function (Vapnik, 2001), the regression problem can be written in the form of convex optimization problem (Smola & Schölkopf, 2004) as

$$\operatorname{Minimize} \frac{1}{2} \left\| \boldsymbol{w}^2 \right\| \tag{4}$$

Subject to 
$$y_i - (\mathbf{x}_i \mathbf{w} + b) \le \varepsilon$$
,  $i = 1, 2, ..., n$ , (5)

$$\left(\boldsymbol{x}_{i}\boldsymbol{w}+\boldsymbol{b}\right)-\boldsymbol{y}_{i}\leq\boldsymbol{\varepsilon}, i=1,2,\ldots,n, \qquad (6)$$

where  $\varepsilon > 0$  is predetermined constant which controls the noise tolerance.

By introducing non negative slack variables  $\xi_i$  and  $\xi_i^*$  (which measures the deviations of training samples outside  $\varepsilon$  insensitive zone), the above optimization problem becomes (Vapnik, 2001):

Min. 
$$\frac{1}{2} \| \mathbf{w} \|^2 + C \sum_{i=1}^n (\xi_i + \xi_i^*)$$
 (7)

Subject to 
$$y_i - (x_i w + b) \le \varepsilon + \xi_i$$
, (8)

$$(\boldsymbol{x}_{i}\boldsymbol{w}+\boldsymbol{b})-\boldsymbol{y}_{i}\leq\boldsymbol{\varepsilon}+\boldsymbol{\xi}_{i}^{*}, \qquad (9)$$

and 
$$\xi_i, \xi_i^* \ge 0, i=1, 2, ..., n,$$
 (10)

where C>0 is the regularization factor which determines (the cost of error) tradeoff between flatness of regression function and amount of deviations outside the  $\varepsilon$  insensitive zone which are tolerated.

Using Lagrange's multipliers method, the dual of above optimization problem can be expressed as (Gunn, 1998),

$$\max \cdot -\frac{1}{2} \sum_{i=1}^{n} \sum_{j=1}^{n} \left( \alpha_{i} - \alpha_{i}^{*} \right) \left( \alpha_{j} - \alpha_{j}^{*} \right) \mathbf{x}_{i} \mathbf{x}_{j}^{*} - \varepsilon \sum_{i=1}^{n} \left( \alpha_{i} - \alpha_{i}^{*} \right) + \sum_{i=1}^{n} \left( \alpha_{i} - \alpha_{i}^{*} \right) \mathbf{y}_{i}$$
(11)

Subject to 
$$\sum_{i=1}^{n} (\alpha_i - \alpha_i^*) = 0$$
, (12)

and 
$$0 \le \alpha_i \le C$$
,  $0 \le \alpha_i^* \le C$ , (13)

where  $\alpha_i$  and  $\alpha_i^*$ , i=1,2,...,n are Lagrange's multipliers that act as forces pushing the predictions towards the target value  $y_i$ .

Above quadratic programming problem can be solved for obtaining the values of  $\alpha_i$  and  $\alpha_i^*$ . Using Karush-Kuhn-Tucker conditions (Smola & Schölkopf, 2004) the weight vector is given by

$$\mathbf{w}' = \sum_{i=1}^{n_{nsv}} \left( \alpha_i - \alpha_i^* \right) \mathbf{x}_i \tag{14}$$

and 
$$f(\mathbf{x}) = \sum_{i=1}^{n_{nsv}} (\alpha_i - \alpha_i^*) \mathbf{x}_i \mathbf{x}' + b$$
, (15)

where  $n_{nsv}$  denotes the number of support vectors.

The value of bias b is given by (Gunn, 1998),

$$b = -\frac{1}{2} (\boldsymbol{x}_r + \boldsymbol{x}_s) \boldsymbol{w}$$
(16)

where  $x_i$ , and  $x_s$  are the support vectors (i.e. any input vector which has nonzero value of either  $\alpha_i$  or  $\alpha_i^*$  respectively).

The performance of SVR strongly depends on proper setting of regularization parameter (C) and the value of  $\varepsilon$ . Such parameters are called as meta parameters. The values of meta parameters C and  $\varepsilon$  are not known in advance and must be obtained from the training data.

# Subset selection using SSp statistic

Consider a regression model in Equation (1) as a full model. Partition the X matrix and vector  $\boldsymbol{\beta}$  as

$$\mathbf{X} = \left[\mathbf{X}_{(1)} : \mathbf{X}_{(2)}\right] \text{ and } \boldsymbol{\beta} = \left[\boldsymbol{\beta}_{(1)} : \boldsymbol{\beta}_{(2)}\right]',$$

where  $\mathbf{X}_{(1)}$  is an  $n \times p$  matrix of observations on p-1 predictor variables with 1's in the first column and  $\mathbf{X}_{(2)}$  is an  $n \times (k-p)$  matrix of observations on remaining (k-p) predictor variables.  $\boldsymbol{\beta}_{(1)} = (\beta_0, \beta_p, \beta_{2^n}, \beta_{p-1})'$  is  $p \times 1$  vector of regression parameters corresponding to (p-1) predictor variables and  $\boldsymbol{\beta}_{(2)}$  is  $(k-p \times 1 \text{ vector})$ of regression parameters corresponding to remaining (k-p) predictor variables. In these notations the full model becomes

$$Y = X_{(1)}\beta_{(1)} + X_{(2)}\beta_{(2)} + e$$
(17)

A subset model based on a (p-1) predictors is given by

$$Y = X_{(1)}\beta_{(1)} + e$$
 (18)

Similarly, for SVR partition the weight vector w and write the sub model as

$$Y = b_{(1)} + X_{(1)} w_{(1)} + e$$
(19)

where **X**<sub>(1)</sub> is an  $n \times (p-1)$  matrix of the observations on (p-1) predictors and  $w_{(1)}$  is a  $(p-1) \times 1$  vector of the regression coefficients based on the fitted sub model.

For a subset model, the dual of optimization problem can be expressed as

$$\max . -\frac{1}{2} \sum_{i=1}^{n} \sum_{j=1}^{n} \left( \alpha_{i(1)} - \alpha_{i(1)}^{*} \right) \left( \alpha_{j(1)} - \alpha_{j(1)}^{*} \right) \mathbf{x}_{i(1)} \mathbf{x}_{j}^{'} - \varepsilon \sum_{i=1}^{n} \left( \alpha_{i(1)} - \alpha_{i(1)}^{*} \right) + i = Inai(1) - ai(1) * yi$$
(20)

Subject to: 
$$\sum_{i=1}^{n} \left( \alpha_{i(1)} - \alpha_{i(1)}^{*} \right) = 0$$
 (21)

and  $0 \le \alpha_{i(1)} \le C$ ,  $0 \le \alpha_{i(1)}^* \le C$ , where  $\alpha_{i(1)}$  and  $\alpha_{i(1)}^*$ ,  $i=1,2,\ldots,n$  are Lagrange's multipliers. The weight vector is given by

$$\mathbf{w}_{(1)}' = \sum_{i=1}^{n_{nsv}} \left( \alpha_{i(1)} - \alpha_{i(1)}^* \right) \mathbf{x}_{i(1)}$$
(22)

and

$$f(\mathbf{x}) = \sum_{i=1}^{n_{nsv}} \left( \alpha_{i(1)} - \alpha_{i(1)}^* \right) \mathbf{x}_{i(1)} \mathbf{x}' + b_{(1)}.$$
(23)

The value of bias  $b_{(1)}$  is given by (Gunn-1998),

$$b_{(1)} = -\frac{1}{2} \left( \mathbf{x}_{r(1)} + \mathbf{x}_{s(1)} \right) \mathbf{w}_{(1)}$$
(24)

By obtaining estimates  $\hat{w}$  of w and  $\hat{b}$  of b using SVM, the predicted value of y based on the full model is given by

$$\hat{\mathbf{y}}_{ik} = \mathbf{x}_i \, \hat{\mathbf{w}} + \hat{b} \,, \, i = 1, 2, \dots, n$$
 (25)

The predicted value of y based on sub model is given by

$$\hat{y}_{ip} = \boldsymbol{x}_{i(1)} \, \hat{\boldsymbol{w}}_{(1)} + \hat{b}_{(1)}, \ i = 1, 2, \dots, n \tag{26}$$

where  $\hat{\boldsymbol{w}}_{(1)}$  is the estimator of  $\boldsymbol{w}_{(1)}$ ,  $\hat{b}_{(1)}$  is the estimator of  $b_{(1)}$  and  $\boldsymbol{x}_{i(1)} \in R^{p-1}$ (*i*<sup>th</sup> row of  $\mathbf{X}_{(1)}$  excluding first element 1)

Kashid and Kulkarni (2002) defined the Sp Statistic for subset selection based on predicted values from full and subset models using M-estimator given by

#### SVM-BASED MODIFIED SP FOR SUBSET SELECTION

$$Sp = \sum_{i=1}^{n} \frac{\left(\hat{y}_{ik} - \hat{y}_{ip}\right)^{2}}{\sigma^{2}} - \left(k - 2p\right), \qquad (27)$$

where,  $\sigma^2$  is replaced by its estimate usually based on full model, k and p are number of parameters in full model and subset model, respectively.

The Sp statistic takes into account closeness of predictions obtained from subset model and incorporates the complexity in the form of number of predictors involved in the model. The penalty term doesn't increase with sample size. If the Sp statistic is based on the estimates obtained from SVR, its performance is not good. Hence, a complexity term is added to the criterion so as to increase its ability to identify the correct subset model. The modified Sp statistic is called the SSp statistic and is given by

$$SSp = \sum_{i=1}^{n} \frac{\left(\hat{y}_{ik} - \hat{y}_{ip}\right)^{2}}{\sigma^{2}} - (k - p) + g(n, p), \qquad (28)$$

where error variance  $\sigma^2$  is usually unknown and is to be replaced by its suitable estimate. As discussed in Bozdogan and Haughton (1998) the term g(n, p) is a nonnegative penalty function which increases as the number of parameters increases. In this study, to calculate SSp statistic use SVR estimates which are robust, and thus SSp is also robust.

For g(n, p) = p,  $\beta$  is replaced by its LS estimate then SSp statistic is equivalent to Mallow's Cp statistic, which is given by

Cp = 
$$\sum_{i=1}^{n} \frac{(y_i - \hat{y}_{ip})^2}{\sigma^2} - (n - 2p).$$
 (29)

If g(n,p)=p and the M estimator of  $\beta$  is used in (28), then SSp statistic is equivalent to the Sp statistic.

# Simulation results

The estimation accuracy of SVR strongly depends on selection of meta parameters C and  $\varepsilon$ . When the values of meta parameters C and  $\varepsilon$  are not known they have to be obtained from the data itself. There are several methods available in the literature for selection of C amongst which, the method proposed by Desai and Kashid (2015) uses C = MQD for better performance. In the simulations to perform SVR, this method was used. Take  $\varepsilon = C \times 10^{-6}$  and C = MQD = Max.(|Me - 3QD|, |Me + 3QD|), where Me and QD are median and quartile deviation of y values respectively.

#### A simulation design

The following simulation study is carried out for the performance evaluation of the subset selection methods. The observations on predictor variables are generated from U(0, 1) and are fixed. Observations on error term are generated from N(0,1), students t distribution with 2 d.f., Laplace(0,1), Mixture of Normal  $\{0.2N(5,1) + 0.8N(0,1)\}$ , standard Cauchy and Slash distributions {ratio of N(0,1) and U(0,1)}.

To obtain the observations on response variable following models are used.

Model-I: 
$$Y = \beta_0 + X_1\beta_1 + X_2\beta_2 + X_3\beta_3 + e$$
  
Model-II:  $Y = \beta_0 + X_1\beta_1 + X_2\beta_2 + X_3\beta_3 + X_4\beta_4 + e$ 

For model-I, we consider three sets of regression parameters (2, 5, 0, 0), (2, 5, 4, 0) and (2, 5, 4, 8), and call them as Model-IA, Model-IB and Model-IC respectively. For model-II, take four sets of regression parameters (2, 0, 6, 0, 0), (2, 9, 6, 0, 0), (2, 9, 0, 4, 8) and (2, 9, 6, 4, 8), and call them as Model-IIA, Model-IIB, Model-IIC and Model-IID, respectively.

For each sample size 20, 30, 50 and 100 the experiment is repeated 1000 times. To evaluate the performance of different methods, the probabilities of selecting the optimal and correct models are calculated for various combinations of error distributions and sample sizes. SSp statistic with six different penalty functions is used. The SSp statistic using the penalty functions  $g_1(n,p) = p$ ,  $g_2(n,p) = 2p$ ,  $g_3(n,p) = p\sqrt{n}$ ,  $g_4(n,p) = p(\sqrt{n}+1)$ ,  $g_5(n,p) = plog(n)$  and  $g_6(n,p) = p(\log(n)+1)$  is called as SSp1, SSp2, SSp3, SSp4, SSp5 and SSp6 respectively.

In regression true value of variance is often known. If it is not known, it has to be estimated from the data itself. The estimation of variance plays an important role in regression. From the available estimators of  $\sigma^2$ , use  $\hat{\sigma}_1^2 = (1.4826 \text{MAD})^2$ to calculate Sp, where  $MAD = Median |r_i - Median(r_i)|$  and  $r_i = y_i - \hat{y}_i$  (Birkes & Dodge, 1993). The performance of  $\hat{\sigma}_1^2$  estimator works well when the M-estimator is used for obtaining regression coefficients (Kashid & Kulkarni, 2002). If SVR is used for the prediction obtain  $\hat{\sigma}_1^2$ , then  $\hat{\sigma}_1^2$  will under-estimates the true value of  $\sigma^2$ . This is verified in the simulation study. To improve the performance, define  $\hat{\sigma}_2^2$ using SVR estimates as,

$$\hat{\sigma}_{2}^{2} = \left[1.4826 \times \left(0.8 + \frac{k}{n}\right) \times \mathbb{P}_{60}\left\{abs\left(r_{i}\right)\right\}\right]^{2}$$
(30)

where  $P_{60}$  is 60<sup>th</sup> percentile of absolute residuals, k and p are the number of parameters of the full and subset model respectively. The performance of  $\hat{\sigma}_2^2$  using SVR estimates is verified in the simulation study (not shown here). In the simulation study for calculation of SSp statistics, use  $\hat{\sigma}_2^2$ . For clarity of notations, call Sp1 for Sp statistic when estimator of variance is  $\hat{\sigma}_1^2$  and Sp2 for Sp statistic when estimator of variance  $\hat{\sigma}_2^2$  is used.

The M-estimator is calculated using Huber loss function (Huber, 1981) with bending point 1.345. Also, for comparison purposes, consider the Robust Akaike's Information Criterion (RAIC) (Ronchetti, 1985) and Robust Bayes Information Criterion (RBIC) (Baierl et al., 2007). The RAIC is defined as

RAIC = 
$$2\sum_{i=1}^{n} \rho(r_{i;p}) + 2p$$
 (31)

where  $\rho(\cdot)$  is Huber function and  $r_{i;p} = (y_i - \boldsymbol{x}_{i(1)}^* \boldsymbol{\tilde{\beta}}_{(1)})/\hat{\sigma}$ ,  $\hat{\sigma}$  is some robust estimate of  $\sigma$ ,  $\boldsymbol{\tilde{\beta}}_{(1)}$  is the M-estimator of  $\boldsymbol{\beta}_{(1)}$  and  $\boldsymbol{x}_{i(1)}^* \in R^p$  (*i*<sup>th</sup> row of  $\mathbf{X}_{(1)}$ ). RBIC is defined as

RBIC = 
$$nlog \sum_{i=1}^{n} \rho(y_i^{(s)} - \mathbf{x}_{i(1)}^* \tilde{\boldsymbol{\beta}}_{(1)}) + plog(n)$$
 (32)

where,  $y_i^{(s)}$  is standardized observation obtained by subtracting the median and dividing by 1.486MAD of y values,  $\tilde{\beta}_{(1)}$  is the M-estimator of  $\beta_{(1)}$ . Due to space constraint some of the simulation results are summarized in the Tables 1 to 3 and Figures 1 to 4.

From Tables 1–3 and Figures 1–4 we observed the following:

For model IA, predictors  $X_2$  and  $X_3$  are redundant, performance of SSp3 and SSp4 is better than RAIC, Sp1 and Sp2 and it is compatible with RBIC. For model IB, predictor  $X_3$  is redundant, performance of SSp3 and SSp4 is better than RAIC, RBIC, Sp1 and Sp2 (except for Slash distribution). For sample size n = 100 and Cauchy and Slash errors, the performance of SSp3, SSp4, SSp5 and SSp6 is better than RAIC, RBIC, RBIC, Sp1 and Sp2. For Model IC which is full as well as optimal model, performance of SSp1 is compatible with Sp1, Sp2, RAIC and RBIC for large sample size.

For Model IIA, performance of SSp3, SSp4, SSp5 and SSp6 is better than SSp1, SSp2, RAIC, Sp1 and Sp2. Also it is compatible with RBIC. The performance of SSp2 is better than SSp1, SSp4 is better than SSp3, and SSp6 is better than SSp5 (except for large sample). For Models IIB and IIC, the performance of SSp3, SSp4, SSp5 and SSp6 is better than SSp1, SSp2, RAIC, Sp1 and Sp2, and is compatible with RBIC. Model IID is full model and optimal model. For this model, performance of SSp1 is compatible with Sp1, Sp2, RAIC and RBIC for large sample size. For

SSp with New S.D. and penalty												
Error	S.S.		SSp1	SSp2	SSp3	SSp4	SSp5	SSp6	Sp1	Sp2	RAIC	RBIC
	20	Optimal	0.633	0.706	0.766	0.759	0.739	0.766	0.757	0.820	0.745	0.87
N(0, 1)	20	Correct	0.349	0.261	0.134	0.104	0.200	0.150	0.239	0.172	0.253	0.12
	30	Optimal	0.647	0.773	0.913	0.932	0.857	0.889	0.775	0.824	0.797	0.91
	50	Correct	0.353	0.226	0.079	0.056	0.141	0.108	0.225	0.176	0.203	0.08
N(O	50	Optimal	0.662	0.772	0.947	0.960	0.882	0.905	0.802	0.827	0.806	0.94
	50	Correct	0.338	0.228	0.053	0.040	0.118	0.095	0.198	0.173	0.194	0.05
	100	Optimal	0.652	0.765	0.976	0.988	0.901	0.927	0.817	0.832	0.822	0.97
	_ 100	Correct	0.348	0.235	0.024	0.012	0.099	0.073	0.183	0.168	0.178	0.02
	20	Optimal	0.589	0.633	0.579	0.560	0.628	0.595	0.653	0.714	0.682	0.70
	20	Correct	0.306	0.208	0.080	0.058	0.137	0.095	0.279	0.184	0.261	0.06
	30	Optimal	0.670	0.753	0.835	0.825	0.813	0.833	0.722	0.784	0.721	0.81
2	30	Correct	0.312	0.213	0.060	0.039	0.131	0.082	0.269	0.204	0.259	0.03
t-2	50	Optimal	0.690	0.786	0.944	0.949	0.891	0.924	0.742	0.794	0.753	0.95
	50	Correct	0.308	0.211	0.042	0.032	0.105	0.068	0.258	0.206	0.246	0.01
		Optimal	0.712	0.814	0.985	0.990	0.932	0.950	0.752	0.792	0.748	0.99
	100	Correct	0.288	0.186	0.015	0.010	0.068	0.050	0.248	0.208	0.252	0.00
						S.D. and						
Irror	S.5	S.Model	SSp1	SSp2	SSp3	SSp4	SSp5	SSp6	Sp1	Sp2	RAIC	RBIC
	20	Optimal	0.621	0.682	0.702	0.692	0.712	0.710	0.669	0.740	0.704	0.87
	20	Correct	0.340	0.243	0.112	0.081	0.172	0.131	0.305	0.224	0.285	0.08
a)	30	Optimal	0.680	0.780	0.894	0.888	0.857	0.885	0.705	0.783	0.748	0.92
ac	30	Correct	0.316	0.209	0.065	0.055	0.126	0.090	0.295	0.216	0.251	0.05
Laplace	50	Optimal	0.698	0.808	0.975	0.978	0.920	0.950	0.763	0.815	0.768	0.97
	50	Correct	0.301	0.191	0.024	0.020	0.079	0.049	0.237	0.185	0.232	0.02
	100	Optimal	0.752	0.857	0.992	0.994	0.945	0.961	0.778	0.804	0.775	0.97
	100	Correct	0.248	0.143	0.008	0.006	0.055	0.039	0.222	0.196	0.225	0.02
	-	Optimal	0.618	0.703	0.780	0.773	0.740	0.778	0.725	0.807	0.769	0.87
	20	Correct	0.364	0.260	0.119	0.087	0.194	0.134	0.271	0.187	0.231	0.11
		Optimal	0.649	0.760	0.917	0.936	0.854	0.887	0.773	0.817	0.787	0.90
Mixture	30	Correct	0.351	0.240	0.080	0.059	0.144	0.111	0.227	0.183	0.213	0.09
lixt		Optimal	0.640	0.769	0.948	0.958	0.880	0.906	0.796	0.837	0.812	0.94
2	50	Correct	0.360	0.231	0.052	0.042	0.120	0.094	0.204	0.163	0.188	0.05
		Optimal	0.670	0.778	0.978	0.980	0.902	0.935	0.829	0.848	0.809	0.95
	100	Correct	0.330	0.222	0.022	0.020	0.098	0.065	0.171	0.152	0.191	0.04
		Optimal	0.545	0.529	0.429	0.385	0.492	0.452	0.539	0.565	0.517	0.24
	20	Correct	0.190	0.123	0.042	0.027	0.081	0.050	0.230	0.147	0.245	0.02
		Optimal	0.664	0.722	0.617	0.573	0.712	0.671	0.619	0.678	0.609	0.25
chy	30	Correct	0.255	0.138	0.028	0.020	0.065	0.044	0.275	0.193	0.242	0.00
		Optimal	0.729	0.802	0.862	0.840	0.881	0.887	0.722	0.780	0.685	0.38
Cal	50	Correct	0.260	0.171	0.020	0.013	0.072	0.046	0.260	0.192	0.287	0.00
		Optimal	0.200 0.762	0.171 0.857	0.020 0.988	0.013 0.986	0.072 0.956	0.040 0.968	0.200	0.192	0.207	
	100	Correct	0.238	0.142	0.005	0.004	0.041	0.029	0.287		0.730	0.58
	-									0.226		0.00
	20	Optimal Corroct	0.374	0.311	0.192	0.159	0.252	0.210	0.399	0.395	0.398	0.11
		Correct	0.150	0.085	0.025	0.016	0.054	0.031	0.204	0.106	0.147	0.00
F	30	Optimal	0.509	0.496	0.311	0.265	0.436	0.373	0.540	0.544	0.487	0.12
Slash		Correct	0.203	0.106	0.014	0.009	0.042	0.024	0.199	0.140	0.192	0.00
Sla	50	Optimal	0.669	0.704	0.510	0.461	0.659	0.615	0.692	0.718	0.643	0.16
	50	Correct	0.243	0.150	0.008	0.002	0.053	0.032	0.222	0.170	0.240	0.00
		Conect	0.2.0									
	100	Optimal	0.723	0.809	0.841	0.818	0.902	0.909	0.732	0.763	0.719	0.28

Table 1. Probabilities of selecting Optimal and Over fitted (Correct) models for Model I B

SSp with New S.D. and penalty												
Error	S.S.	Model	SSp1	SSp2	SSp3	SSp4	SSp5	SSp6	Sp1	Sp2	RAIC	RBIC
	20	Optimal	0.549	0.647	0.741	0.744	0.698	0.729	0.694	0.779	0.735	0.867
	20	Correct	0.436	0.322	0.170	0.137	0.247	0.190	0.303	0.214	0.262	0.125
-	30	Optimal	0.597	0.688	0.872	0.895	0.788	0.832	0.752	0.812	0.774	0.908
N(0, 1)	30	Correct	0.403	0.312	0.124	0.098	0.211	0.165	0.248	0.188	0.226	0.092
N(O	50	Optimal	0.609	0.722	0.935	0.948	0.850	0.891	0.792	0.836	0.822	0.959
_	50	Correct	0.391	0.278	0.064	0.050	0.149	0.108	0.208	0.164	0.178	0.041
	100	Optimal	0.612	0.734	0.982	0.986	0.882	0.916	0.820	0.848	0.852	0.977
	100	Correct	0.388	0.266	0.018	0.014	0.118	0.084	0.180	0.152	0.148	0.023
	20	Optimal	0.581	0.638	0.607	0.576	0.640	0.622	0.650	0.725	0.666	0.692
	20	Correct	0.334	0.215	0.089	0.064	0.152	0.106	0.286	0.174	0.267	0.079
	30	Optimal	0.635	0.725	0.867	0.864	0.813	0.856	0.730	0.815	0.731	0.873
t-2	30	Correct	0.353	0.254	0.077	0.060	0.154	0.102	0.269	0.179	0.260	0.052
4	50	Optimal	0.672	0.772	0.934	0.928	0.875	0.898	0.743	0.792	0.751	0.968
	50	Correct	0.325	0.224	0.033	0.026	0.113	0.086	0.253	0.204	0.249	0.022
	100	Optimal	0.662	0.788	0.988	0.992	0.912	0.942	0.742	0.804	0.773	0.990
	100	Correct	0.338	0.212	0.012	0.008	0.088	0.058	0.258	0.196	0.227	0.008
	20	Optimal	0.607	0.670	0.714	0.698	0.711	0.719	0.669	0.756	0.671	0.824
	20	Correct	0.354	0.259	0.118	0.092	0.180	0.130	0.297	0.189	0.304	0.113
(I)	30	Optimal	0.634	0.758	0.894	0.911	0.840	0.873	0.699	0.788	0.737	0.936
Laplace	30	Correct	0.364	0.240	0.094	0.067	0.156	0.120	0.301	0.209	0.262	0.060
ap	50	Optimal	0.668	0.775	0.956	0.957	0.890	0.923	0.740	0.815	0.756	0.978
	50	Correct	0.332	0.225	0.038	0.031	0.108	0.074	0.260	0.185	0.244	0.022
	100	Optimal	0.718	0.810	0.982	0.990	0.938	0.960	0.762	0.808	0.777	0.979
	100	Correct	0.282	0.190	0.018	0.010	0.062	0.040	0.238	0.192	0.223	0.021
	- 20 30	Optimal	0.561	0.657	0.749	0.760	0.713	0.745	0.701	0.798	0.734	0.879
		Correct	0.425	0.316	0.165	0.124	0.242	0.186	0.293	0.192	0.263	0.116
		Optimal	0.563	0.685	0.868	0.897	0.789	0.833	0.751	0.814	0.797	0.914
Mixture	30	Correct	0.437	0.315	0.130	0.101	0.211	0.166	0.249	0.186	0.203	0.086
Mix	50	Optimal	0.615	0.717	0.930	0.942	0.852	0.885	0.788	0.844	0.803	0.936
_	50	Correct	0.385	0.283	0.068	0.055	0.148	0.115	0.212	0.156	0.197	0.064
	100	Optimal	0.642	0.760	0.974	0.984	0.902	0.924	0.844	0.860	0.836	0.971
	100	Correct	0.358	0.240	0.026	0.016	0.098	0.076	0.156	0.140	0.164	0.029
	20	Optimal	0.506	0.485	0.393	0.349	0.465	0.419	0.535	0.534	0.520	0.245
	20	Correct	0.233	0.152	0.049	0.036	0.098	0.054	0.235	0.142	0.228	0.018
>	30	Optimal	0.673	0.739	0.745	0.710	0.773	0.772	0.685	0.752	0.617	0.366
Cauchy	50	Correct	0.285	0.185	0.040	0.032	0.109	0.067	0.253	0.170	0.297	0.010
Cal	50	Optimal	0.697	0.790	0.808	0.769	0.854	0.849	0.688	0.762	0.693	0.492
0	50	Correct	0.281	0.169	0.018	0.013	0.071	0.046	0.289	0.205	0.295	0.004
	100	Optimal	0.756	0.848	0.982	0.980	0.958	0.966	0.744	0.804	0.700	0.604
	- 100	Correct	0.242	0.150	0.002	0.002	0.040	0.030	0.252	0.192	0.300	0.000
	20	Optimal	0.330	0.281	0.160	0.123	0.229	0.177	0.382	0.345	0.375	0.125
	20	Correct	0.157	0.080	0.030	0.021	0.047	0.033	0.173	0.092	0.197	0.016
	30	Optimal	0.548	0.565	0.448	0.404	0.550	0.505	0.582	0.626	0.531	0.195
Slash	50	Correct	0.278	0.183	0.040	0.025	0.093	0.062	0.244	0.155	0.240	0.007
Sĩ	50	Optimal	0.610	0.673	0.524	0.485	0.657	0.620	0.647	0.691	0.688	0.294
	50	Correct	0.283	0.170	0.023	0.013	0.071	0.049	0.249	0.176	0.260	0.002
	100	Optimal	0.702	0.802	0.828	0.804	0.914	0.916	0.736	0.788	0.724	0.349
	100	Correct	0.294	0.190	0.004	0.004	0.048	0.034	0.254	0.202	0.268	0.000

Table 2. Probabilities of Selecting Optimal and Over fitted (Correct) models for Model II C

Size		5	Sp-(SVR	with New	<sup>,</sup> S.D.) wi	у					
Ś	Error	SSp6	SSp6	SSp6	SSp6	SSp6	SSp6	Sp1	Sp2	RAIC	RBIC
	N(0, 1)	0.991	0.977	0.936	0.919	0.963	0.953	0.997	0.992	1.000	0.999
	t2	0.926	0.886	0.739	0.694	0.836	0.771	0.935	0.899	0.957	0.798
0	L(0, 1)	0.966	0.944	0.871	0.839	0.914	0.888	0.966	0.944	0.986	0.960
20	Mixture	0.987	0.984	0.941	0.915	0.973	0.947	0.994	0.990	1.000	0.999
	Cauchy	0.777	0.669	0.499	0.433	0.588	0.530	0.758	0.664	0.753	0.266
	Slash	0.510	0.401	0.208	0.159	0.310	0.226	0.527	0.420	0.572	0.147
	N(0, 1)	1.000	0.997	0.984	0.972	0.992	0.987	1.000	1.000	1.000	1.000
	t2	0.971	0.957	0.857	0.805	0.924	0.896	0.975	0.962	0.991	0.910
30	L(0, 1)	0.997	0.990	0.939	0.919	0.968	0.952	0.995	0.992	0.999	0.994
с	Mixture	0.997	0.995	0.976	0.961	0.989	0.984	1.000	0.999	1.000	1.000
	Cauchy	0.884	0.827	0.593	0.536	0.753	0.678	0.852	0.808	0.884	0.308
	Slash	0.710	0.607	0.308	0.248	0.475	0.389	0.665	0.584	0.754	0.167
	N(0, 1)	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000
	t2	0.999	0.999	0.991	0.988	0.997	0.997	0.999	0.999	0.999	0.973
50	L(0, 1)	1.000	1.000	0.999	0.998	0.999	0.999	1.000	1.000	1.000	1.000
Ŋ	Mixture	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000
	Cauchy	0.984	0.978	0.898	0.874	0.960	0.946	0.985	0.978	0.967	0.365
	Slash	0.939	0.897	0.621	0.562	0.802	0.744	0.928	0.901	0.882	0.144
	N(0, 1)	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000
	t2	0.999	0.999	0.999	0.999	0.999	0.999	1.000	1.000	1.000	0.999
00	L(0, 1)	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000
1	Mixture	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000	1.000
	Cauchy	0.996	0.996	0.981	0.978	0.991	0.990	1.000	1.000	0.997	0.469
	Slash	0.994	0.991	0.857	0.824	0.959	0.945	0.994	0.994	0.977	0.237

Table 3. Probabilities of selecting Optimal (and Correct model) for Model II D

large samples almost all criteria select optimal model with probability 1 except for Cauchy and Slash errors. For small samples (n = 20) the performance of RBIC is greater than SSp so, for a small sample size it is suggested to use RBIC.

In the simulation, for every model there are six different values of probabilities of selecting optimal model corresponding to six different error distributions. For sample size n = 100, obtain the average and standard deviation of these six probabilities for each model. The results are reported in Table 4.

From the summary statistics it is observed that when full model is an optimal model then SSp1 and SSp2 perform better than all other statistics considered in this simulation. When full model is not an optimal model then SSp3 and SSp4 works better than others.

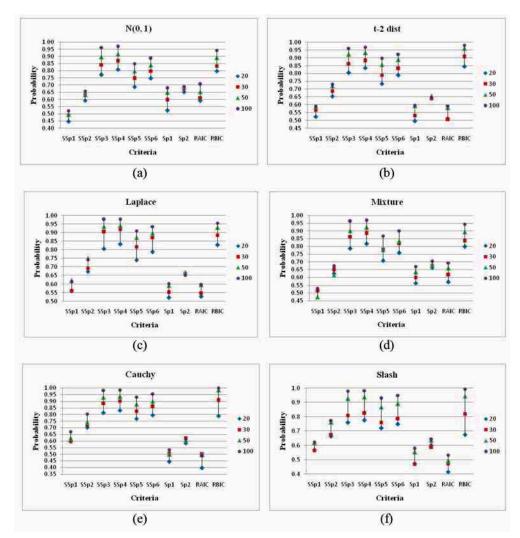


Figure 1. Probabilities of selecting optimal model : Model-I A

#### **Real Data Application**

**Example 1:** To observe the performance of various criteria, consider Brownlee's stack loss data (Hand et al., 1994, pp 156) which contains observations from 21 days operation of a plant for the oxidation of ammonia as a stage in the production of nitric acid. The predictor variables are  $X_1$  = air flow,  $X_2$  = cooling water inlet temperature (° C),  $X_3$  = acid concentration (%) and the response variable Y = stack loss. Stack loss is the percentage of the ingoing ammonia that escapes unabsorbed.

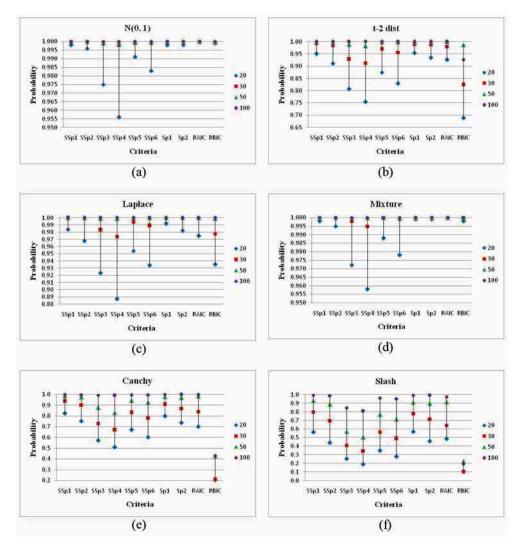


Figure 2. Probabilities of selecting optimal model: Model-I C

As mentioned in Montgomery et al. (2006, p. 396) observation no. 21 is an influential observation because it has a standardized residual of -2.64 from LS fit (Rousseuw & Leroy, 1987, p. 226/227; Rousseuw & van Zomeren, 1990). For a single outlier, replace observation 21 by 7.5 instead of original value 1.5. As a result, standardized residual corresponding to observation no. 21 becomes 4.01 from LS fit, which indicates that observation no. 21 is a potential outlier. Similarly, for two outliers, replace observation 21 by 7.5 and observation14 by 6 instead of original value 1.2, the corresponding standardized residuals from LS fit are 2.63 and 2.56 respectively, which indicates that observations 21 and 14 are outliers.

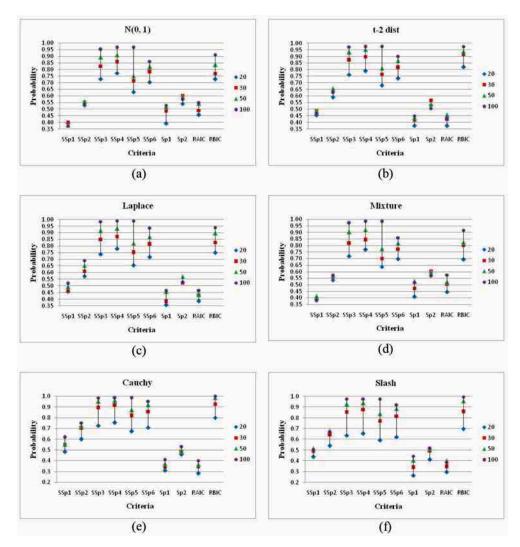


Figure 3. Probabilities of selecting optimal model : Model-II A

For original and outlier data, calculate the SSp statistic with various penalties and  $\hat{\sigma}_2^2$ , Sp statistic using SVR (with C = MQD and  $\varepsilon$  = C ×10<sup>-6</sup>) estimates and Mestimator, also Cp statistic using LS. Sp1 refers to the Sp statistic using M estimator and  $\hat{\sigma}_1^2$  and Sp2 for Sp statistic using M estimator and  $\hat{\sigma}_2^2$ . Also, for comparison purpose, calculate RAIC and RBIC for the original and outlier data. The results are presented in the Table 5.

For the original data, all statistics Sp, Cp, SSp with all penalties, Sp1, Sp2, RAIC and RBIC select the subset  $\{X_1, X_2\}$ . For one outlier, statistics Sp, SSp with all

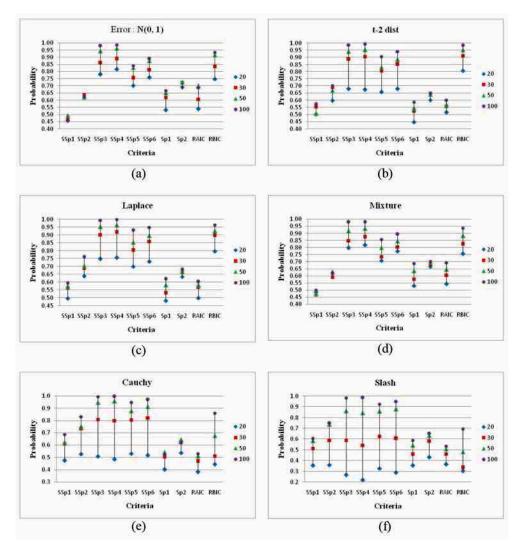


Figure 4. Probabilities of selecting optimal model : Model-II B

penalties and RAIC select the same subset  $\{X_1, X_2\}$  but Cp, Sp2 and RBIC select wrong subset  $\{X_1\}$ . For two outliers, statistics SSp with all penalties select the same subset  $\{X_1, X_2\}$ , while others select wrong subsets. In particular, Sp and Cp select subset  $\{X_1, X_2, X_3\}$  and Sp1, Sp2, RAIC and RBIC select subset  $\{X_1\}$ .

**Example 2:** Consider the wine quality data from Montgomery et al. (2006, Table B.14, p. 578), which contains 38 observations on response variable wine quality (Y) based on five predictor variables as clarity  $(X_1)$ , aroma  $(X_2)$ , body  $(X_3)$ , flavor  $(X_4)$ 

Model	Statistic	SSp1	SSp2	SSp3	SSp4	SSp5	SSp6	Sp1	Sp2	RAIC	RBIC
Model IA	Average	0.591	0.730	0.971	0.975	0.898	0.925	0.610	0.658	0.602	0.970
	S.D.	0.058	0.057	0.009	0.007	0.033	0.027	0.056	0.033	0.086	0.025
Model IB	Average	0.712	0.813	0.960	0.959	0.923	0.942	0.770	0.802	0.767	0.793
	S.D.	0.044	0.039	0.059	0.069	0.025	0.022	0.046	0.033	0.042	0.294
Model IC	Average	0.999	0.998	0.973	0.967	0.993	0.992	0.999	0.998	0.996	0.759
	S.D.	0.002	0.004	0.062	0.077	0.015	0.019	0.002	0.004	0.010	0.354
	Average	0.476	0.641	0.973	0.979	0.979	0.905	0.469	0.538	0.467	0.956
Model IIA	S.D.	0.093	0.079	0.011	0.009	0.009	0.038	0.045	0.028	0.079	0.039
Model IIB	Average	0.569	0.715	0.987	0.990	0.901	0.933	0.612	0.666	0.608	0.896
	S.D.	0.082	0.081	0.006	0.007	0.042	0.033	0.059	0.031	0.072	0.107
Madal IIC	Average	0.682	0.790	0.956	0.956	0.918	0.937	0.775	0.819	0.777	0.812
Model IIC	S.D.	0.053	0.040	0.063	0.075	0.027	0.022	0.046	0.028	0.060	0.272
	Average	0.998	0.998	0.973	0.967	0.992	0.989	0.999	0.999	0.996	0.784
IVIOGEI IID	S.D.	0.003	0.004	0.057	0.071	0.016	0.022	0.002	0.002	0.009	0.342
Model IID	S.D.	0.003	0.004	0.057	0.071	0.016	0.022	0.002	0.002	0.009	0.342

 Table 4. Summary Statistics for probability of selecting optimal model for sample size

 n=100

and oakiness ( $X_5$ ). Test this data for outliers and multicollinearity using Minitab. This data contains only one influential observation, which is observation 20 that has a standardized residual of 2.74. Apply the procedure described in Example 1 for subset selection to this original data. Use all the considered criteria SSp with all penalties, Cp, Sp, Sp1, Sp2, RAIC and RBIC, and select the subset { $X_7$ ,  $X_4$ ,  $X_5$ }.

For a single outlier, replace observation 20 by 39.5 instead of original value 7.9, as a result its standardized residual become 5.51, which indicates that observation 20 is a potential outlier. Apply the same procedure for this outlier data. Using the criteria SSp2, SSp3, SSp4, SSp5 and SSp6, select the subset  $\{X_2, X_4, X_5\}$ . The criteria Sp1, Sp2 and RAIC select a different subset  $\{X_1, X_3, X_4, X_5\}$  and RBIC selects the subset  $\{X_4\}$ .

# Discussion

The modified Sp criterion was used for subset selection in regression in the presence of outliers and/or error distribution is non normal. Implementation of modified Sp criterion requires a penalty term. The choices for the penalty terms are not limited to those mentioned in this paper. A more suitable penalty can give superior performance than listed here. The proposed modification makes the criterion free from assumption of the distribution of errors and is even mitigates the effect of outliers. The simulation results confirm these findings.

Sp (SVM)		Cp .	SSp							Sp2	RAIC	RBIC
Sul	(SVM) Cp	00 -	SSp1	SSp2	SSp3	SSp4	SSp5	SSp6	Sp1	067		1.2.0
Stack L	oss data	a ( Origina	al)									
1*	6.11	13.34	9.37	11.37	16.54	18.54	18.54	15.46	11.95	8.74	42.91	68.42
2	118.77	28.93	182.28	184.28	189.44	191.44	191.44	188.37	65.20	47.69	73.78	80.69
3	290.65	148.26	446.06	448.06	453.22	455.22	455.22	452.15	265.80	194.43	148.05	95.91
12	2.63	2.95	2.97	5.97	13.71	16.71	16.71	12.10	3.44	3.05	34.07	64.60
13	7.51	14.39	10.46	13.46	21.21	24.21	24.21	19.59	12.82	9.91	43.94	70.93
23	125.65	30.16	191.77	194.77	202.52	205.52	205.52	200.90	64.15	47.46	74.41	83.31
123	4.00	4.00	4.00	8.00	18.33	22.33	22.33	16.18	4.40	4.30	34.79	66.67
One Outlier												
1	5.52	1.72	7.99	9.99	15.15	17.15	17.15	14.08	4.32	2.58	114.03	90.25
2	151.59	18.67	219.53	221.53	226.70	228.70	228.70	225.62	136.70	81.69	201.02	102.48
3	337.67	21.72	489.03	491.03	496.20	498.20	498.20	495.12	375.53	224.42	276.59	109.30
12	2.35	2.05	2.51	5.51	13.26	16.26	16.26	11.64	2.98	2.59	113.86	92.88
13	9.07	3.67	12.24	15.24	22.98	25.98	25.98	21.37	5.45	4.06	115.51	93.19
23	158.05	17.87	227.99	230.99	238.74	241.74	241.74	237.12	129.06	77.93	198.61	105.05
123	4.00	4.00	4.00	8.00	18.33	22.33	22.33	16.18	4.09	4.05	115.27	95.81
Τωο Οι	utliers											
1	4.02	4.88	10.60	12.60	17.76	19.76	19.76	16.69	0.87	0.72	132.12	93.45
2	14.04	16.59	225.06	227.06	232.22	234.22	234.22	231.15	82.33	68.48	194.30	101.75
3	8.26	10.22	472.81	474.81	479.98	481.98	481.98	478.90	172.73	143.67	229.06	105.28
12	4.13	4.21	3.16	6.16	13.91	16.91	16.91	12.29	2.50	2.42	134.02	96.47
13	3.83	4.66	23.81	26.81	34.56	37.56	37.56	32.94	3.06	2.88	133.38	96.37
23	9.38	11.38	244.32	247.32	255.07	258.07	258.07	253.45	71.93	60.17	188.15	103.88
123	4.00	4.00	4.00	8.00	18.33	22.33	22.33	16.18	4.75	4.63	135.34	99.41

Table 5. Model selection for Brownlee's stack loss data.

\*Note: The symbol ij means the variables X<sub>i</sub> and X<sub>i</sub> are in the model.

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