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Moment Generating Functions of Complementary Exponential-Geometric Distribution Based on k-th Lower Record Values

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Cover Page Footnote

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The complementary exponential-geometric (CEG) distribution is a useful model for analyzing lifetime data. For this distribution, some recurrence relations satisfied by marginal and joint moment generating functions of k -th lower record values were established. They enable the computation of the means, variances, and covariances of k -th lower record values for all sample sizes in a simple and efficient recursive manner. Means, variances, and covariances of lower record values were tabulated from samples of sizes up to 10 for various values of the parameters.

Keywords: k -th lower record value, marginal and joint moment generating function, recurrence relations, complementary exponential-geometric distribution

Introduction

Record values and record times are of interest. For example, a meteorologist may deal with upper and lower record temperatures and precipitation levels, or a seismologist may be interested in earthquakes of record magnitude. Several attempts may be made in a sporting event to achieve a record, and records are made only when the attempt is a success. Usually, data on attempts to break records are not available.

Chandler (1952) studied the distributions of lower records, record times, and inter-record times for identically independently distributed (iid) sequences of random variables. This model takes a certain dependence structure into consideration. The life-length distribution of the components in the system may change after each failure of the components. For this type of model, consider the lower record statistics. If various voltages of equipment are considered, only the voltages less than the previous one can be recorded. These recorded voltages are the lower record value sequence. The related literature on record-breaking data and its distributional properties is vast (see, for instance, Chandler, 1952; Resnick, 1973; Shorrock, 1973; Glick, 1978; Nevzorov, 1987; Ahsanullah, 1995; Balakrishnan & Ahsanullah, 1993, 1994; Grunzień & Szynal, 1997; Arnold, Balakrishnan, & Nagaraja, 1992, 1998; Pawlas & Szynal, 1999; Kumar, 2012, 2015, 2016; Kumar & Kulshrestha, 2013).

A random variable X is said to have complementary exponential-geometric distribution (see Louzada, Roman, & Cancho, 2011) if its pdf is of the form

$$f(x) = \frac{\theta e^{-x/\beta}}{\beta [e^{-x/\beta} (1-\theta) + \theta]^2}, \quad x > 0, \beta > 0, 0 < \theta < 1 \quad (1)$$

and the corresponding cdf is

$$F(x) = \frac{\theta(1 - e^{-x/\beta})}{[e^{-x/\beta} (1-\theta) + \theta]}, \quad x > 0, \beta > 0, 0 < \theta < 1 \quad (2)$$

Note that the complementary exponential-geometric distribution defined in (1) can expressed as

$$f(x) = \begin{cases} -(1-\theta)[1-F(x)]^2 + [1-F(x)] & (*) \\ -(1-\theta)[F(x)]^2 + (1-2\theta)[F(x)] + \theta & (**) \end{cases}$$

Here β is the scale parameter and θ is the shape parameter. The distribution exhibits decreasing and unimodal pdfs and increasing hazard rates.

The concept of recurrence is useful in reducing the number of operations necessary to obtain a general form for the function under consideration. It is also

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used in characterizing the distributions, permitting the identification of population distribution from the properties of the sample.

One of the motivations here is the relative paucity of work concerning marginal and joint moment generating functions of k -th lower record values. The objective is to establish an explicit expression and recurrence relations for marginal and joint moment generating functions of k -th lower record values in a simple recursive manner. The so-obtained relationships enable us to compute all the moments of lower record values through using some mathematical software (e.g., Mathematica, Maple).

Record Values and Preliminaries

Let $\{X_n, n \geq 1\}$ be a sequence of iid random variables with cdf $F(x)$ and pdf $f(x)$. The j^{th} order statistics of a sample (X_1, X_2, \dots, X_n) is denoted by $X_{j:n}$. For a fixed $k \geq 1$, define the sequence $\{L^{(k)}(n), n \geq 1\}$ of k -th lower record times of X_1, X_2, \dots as follows:

$$L^{(k)}(1) = 1$$

$$L^{(k)}(n+1) = \min \left\{ j > L^{(k)}(n), X_{k:L^{(k)}(n)+k-1} > X_{k:j+k-1} \right\}$$

The sequences $\{Z_n^{(k)}, n \geq 1\}$ with $Z_n^{(k)} = X_{k:L^{(k)}(n)+k-1}$, $n = 1, 2, \dots$ are called the sequences of k -th lower record values of $\{X_n, n \geq 1\}$. For convenience, take $Z_0^{(k)} = 0$. Note that $k = 1$, $Y_1^{(1)} = X_{L(n)}$, $n \geq 1$, i.e. record values of $\{X_n, n \geq 1\}$. The joint pdf of k -th lower record values $Z_1^{(k)}, Z_2^{(k)}, \dots, Z_n^{(k)}$ can be given as the joint pdf of k -th lower record values of $\{-X_n, n \geq 1\}$ is given by

$$f_{Z_1^{(k)}, Z_2^{(k)}, \dots, Z_n^{(k)}}(z_1, z_2, \dots, z_n) = k^n \left(\prod_{i=1}^{n-1} \frac{f(z_i)}{F(z_i)} \right) [F(z_n)]^{k-1} f(z_n), \quad z_1 > z_2 > \dots > z_n$$

In view of above equation, the marginal pdf of $X_{L(n)}^{(k)}$, $n \geq 1$, is given by

$$f_{X_{L(n)}^{(k)}}(x) = \frac{k^n}{(n-1)!} [-\ln(F(x))]^{n-1} [F(x)]^{k-1} f(x) \quad (3)$$

and the joint pdf of $X_{L(m)}^{(k)}$ and $X_{L(n)}^{(k)}$, $1 \leq m < n$, $n > 2$, is given by

$$\begin{aligned} f_{X_{L(m)}^{(k)}, X_{L(n)}^{(k)}}(x, y) &= \frac{k^n}{(m-1)!(n-m-1)!} \left[-\ln(F(x)) \right]^{m-1} \\ &\times \left[-\ln(F(y)) + \ln(F(x)) \right]^{n-m-1} \left[F(y) \right]^{k-1} \frac{f(x)}{F(x)} f(y) \end{aligned} \quad (4)$$

for $x > y$. Let X_1, X_2, \dots, X_n be a random sample of the complementary exponential-geometric distribution with pdf and cdf as in (1) and (2), respectively, and let $X_{L(1)}, X_{L(2)}, \dots, X_{L(n)}$ be the first n lower record values obtained from this sample. Denote the marginal moment generating functions of $X_{L(n):k}$ by $M_{X_{L(n):k}}(t)$ and its j^{th} derivative by $M_{X_{L(n):k}}^{(j)}(t)$. Similarly, let $M_{X_{L(m):k}, X_{L(n):k}}(t_1, t_2)$ and $M_{X_{L(m):k}, X_{L(n):k}}^{(i,j)}(t_1, t_2)$ denote the joint moment generating functions of $X_{L(m):k}$ and $X_{L(n):k}$ and its $(i, j)^{\text{th}}$ partial derivatives with respect to t_1 and t_2 , respectively.

Relations for Marginal Moment Generating Function

Establish the explicit expression and recurrence relations for marginal moment generating functions of k -th lower record values from complementary exponential-geometric distribution as follows:

Theorem 1. For the distribution given in (3), a positive integer $k \geq 1$, and for $n \geq 1$,

$$M_{X_{L(n):k}}(t) = k^n \sum_{p=0}^{\infty} \sum_{l=0}^{\infty} \frac{\Gamma(\beta t + 1)}{\Gamma(\beta t + 1 - p)} \frac{(p)_l}{\theta^p p!! (k + p + l)^n}, \quad t \neq 0 \quad (5)$$

where

$$(t)_l = \begin{cases} t(t+1) \cdots (t+p-1), & p = 1, 2, \dots \\ 1, & p = 0 \end{cases}$$

Proof. From (3)

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$$\begin{aligned}
 M_{X_{L(n)k}}(t) &= \frac{k^n}{(n-1)!} \int_0^\infty e^{tx} \left[-\ln(F(x)) \right]^{n-1} [F(x)]^{k-1} f(x) dx \\
 &= \frac{k^n}{(n-1)!} \int_0^1 \left[1 + \frac{z}{\theta(1-z)} \right]^{\beta t} z^{k-1} [-\ln z]^{n-1} dz \\
 &= \frac{k^n}{(n-1)!} \sum_{p=0}^{\infty} \frac{\Gamma(\beta t + 1)}{\theta^p p! \Gamma(\beta t + 1 - p)} \int_0^1 z^{p+k-1} (1-z)^{-p} [-\ln z]^{n-1} dz \\
 &= \frac{k^n}{(n-1)!} \sum_{p=0}^{\infty} \sum_{l=0}^{\infty} \frac{\Gamma(\beta t + 1)(p)_l}{\theta^p p! l! \Gamma(\beta t + 1 - p)} \int_0^1 z^{p+k+l-1} [-\ln z]^{n-1} dz \\
 &= \frac{k^n}{(n-1)!} \sum_{p=0}^{\infty} \sum_{l=0}^{\infty} \frac{\Gamma(\beta t + 1)(p)_l}{\theta^p p! l! \Gamma(\beta t + 1 - p)} \int_0^\infty e^{-(p+k+l)u} u^{n-1} du
 \end{aligned}$$

where $z = F(x)$ and $u = -\ln z$. The result follows by the definition of complete gamma distribution.

Remark 1. Setting $k = 1$ in (5), deduce the explicit expression for marginal moment generating functions of lower record values from the complementary exponential-geometric distribution.

Recurrence relations for marginal moment generating functions of k -th lower record values from the cdf (2) can be derived in the following theorem:

Theorem 2. For a positive integer $k > 1$, and for $n \geq 1$ and $j = 1, 2, \dots$,

$$\begin{aligned}
 M_{X_{L(n)k+1}}^{(j)}(t) &= \left(\frac{k+1}{k-1} \right)^{n-1} \frac{\theta}{1-\theta} \left[M_{X_{L(n)k-1}}^{(j)}(t) - M_{X_{L(n-1)k-1}}^{(j)}(t) \right] \\
 &\quad + \left(\frac{k+1}{k} \right)^{n-1} \frac{(1-2\theta)}{1-\theta} \left[M_{X_{L(n)k}}^{(j)}(t) - M_{X_{L(n-1)k}}^{(j)}(t) \right] \\
 &\quad - \frac{(k+1)^{n-1}}{k^n(1-\theta)} \left[t M_{X_{L(n)k}}^{(j)}(t) + j M_{X_{L(n-1)k}}^{(j)}(t) \right] + M_{X_{L(n-1)k+1}}^{(j)}(t)
 \end{aligned} \tag{6}$$

Proof. From (*),

$$\begin{aligned}
 M_{X_{L(n)k}}(t) &= \frac{k^n}{(n-1)!} \int_0^\infty e^{tx} [F(x)]^{k-1} [-\ln(F(x))]^{n-1} f(x) \\
 &\quad \times \left[-(1-\theta)\{1-F(x)\}^2 + \{1-F(x)\} \right] dx \\
 &= -\frac{k^n(1-\theta)}{(n-1)!} [I_1 + I_2 - 2I_3] + \frac{k^n}{(n-1)!} [I_1 - I_3]
 \end{aligned} \tag{7}$$

where

$$\begin{aligned}
 I_1 &= \int_0^\infty e^{tx} [F(x)]^{k-1} [-\ln(F(x))]^{n-1} dx \\
 I_2 &= \int_0^\infty e^{tx} [F(x)]^{k+1} [-\ln(F(x))]^{n-1} dx \\
 I_3 &= \int_0^\infty e^{tx} [F(x)]^k [-\ln(F(x))]^{n-1} dx
 \end{aligned}$$

By integrating by parts,

$$I_1 = \frac{(n-1)!}{t(k-1)^{n-1}} \left[M_{X_{L(n-1)k-1}}(t) - M_{X_{L(n)k-1}}(t) \right] \tag{8}$$

$$I_2 = \frac{(n-1)!}{t(k+1)^{n-1}} \left[M_{X_{L(n-1)k+1}}(t) - M_{X_{L(n)k+1}}(t) \right] \tag{9}$$

and

$$I_3 = \frac{(n-1)!}{tk^{n-1}} \left[M_{X_{L(n-1)k}}(t) - M_{X_{L(n)k}}(t) \right] \tag{10}$$

Combining (7) and (8)-(10),

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$$\begin{aligned}
 M_{X_{L(n)k}}(t) &= \frac{\theta k^n}{t(k-1)^{n-1}} \left[M_{X_{L(n-1)k-1}}(t) - M_{X_{L(n)k-1}}(t) \right] \\
 &\quad - \frac{(1-\theta)k^n}{t(k-1)^{n-1}} \left[M_{X_{L(n-1)k+1}}(t) - M_{X_{L(n)k+1}}(t) \right] \\
 &\quad + \frac{(1-2\theta)k}{t} \left[M_{X_{L(n-1)k}}(t) - M_{X_{L(n)k}}(t) \right]
 \end{aligned} \tag{11}$$

Now, differentiating both the sides of (11) j times with respect to t ,

$$\begin{aligned}
 &\left(\frac{1-\theta}{(k+1)^{n-1}} \right) M_{X_{L(n)k+1}}^{(j)}(t) \\
 &= \frac{\theta}{(k-1)} \left[M_{X_{L(n)k-1}}^{(j)}(t) - M_{X_{L(n-1)k-1}}^{(j)}(t) \right] \\
 &\quad - \frac{1}{k^n} \left[t M_{X_{L(n)k}}^{(j)}(t) + j M_{X_{L(n)k}}^{(j-1)}(t) \right] + \left(\frac{1-\theta}{(k+1)^{n-1}} \right) M_{X_{L(n-1)k+1}}^{(j)}(t) \\
 &\quad + \frac{1-2\theta}{k^{n-1}} \left[M_{X_{L(n)k}}^{(j)}(t) - M_{X_{L(n-1)k}}^{(j)}(t) \right]
 \end{aligned} \tag{12}$$

The proof is complete.

Similarly, by using the relation in equation (***) and proceeding as above, obtain the following result:

Theorem 3. For a positive integer $k > 1$, and for $n \geq 1$ and $j = 1, 2, \dots$,

$$\begin{aligned}
 M_{X_{L(n)k+1}}^{(j)}(t) &= \frac{(k+1)^{n-1}}{k^n(1-\theta)} \left[j M_{X_{L(n)k}}^{(j-1)}(t) + t M_{X_{L(n)k}}^{(j)}(t) \right] + M_{X_{L(n-1)k+1}}^{(j)}(t) \\
 &\quad - \frac{(1-2\theta)}{(1-\theta)} \left(\frac{k+1}{k} \right)^{n-1} \left[M_{X_{L(n-1)k}}^{(j)}(t) - t M_{X_{L(n)k}}^{(j)}(t) \right] \\
 &\quad - \frac{\theta}{(1-\theta)} \left(\frac{k+1}{k-1} \right)^{n-1} \left[M_{X_{L(n-1)k-1}}^{(j)}(t) - M_{X_{L(n)k-1}}^{(j)}(t) \right]
 \end{aligned} \tag{13}$$

Proof. Similar to the proof of Theorem 2.

By differentiating both sides of equations (6) and (13) with respect to t and then setting $t = 0$, we can obtain the recurrence relations for single moment of k -th lower record values from the complementary exponential-geometric distribution.

$$\begin{aligned} \mathbb{E}\left[X_{L(n);k+1}^{(j)}\right] &= \left(\frac{(k+1)^{n-1}}{1-\theta}\right) \left\{ \frac{\theta}{k-1} \mathbb{E}\left[X_{L(n);k-1}^{(j)}\right] - \mathbb{E}\left[X_{L(n-1);k-1}^{(j)}\right] \right\} \\ &\quad + \frac{(1-2\theta)}{k^{n-1}} \left\{ \mathbb{E}\left[X_{L(n);k}^{(j)}\right] - \mathbb{E}\left[X_{L(n-1);k}^{(j)}\right] \right\} \\ &\quad - \frac{j}{k^n} \mathbb{E}\left[X_{L(n);k}^{(j-1)}\right] + \frac{(1-\theta)}{(k+1)^{n-1}} \mathbb{E}\left[X_{L(n-1);k+1}^{(j)}\right] \end{aligned}$$

and

$$\begin{aligned} \mathbb{E}\left[X_{L(n);k+1}^{(j)}\right] &= \frac{j(k+1)^{n-1}}{k^n(1-\theta)} \mathbb{E}\left[X_{L(n);k}^{(j-1)}\right] - \left(\frac{1-2\theta}{1-\theta}\right) \left(\frac{k+1}{k}\right)^{n-1} \left\{ \mathbb{E}\left[X_{L(n-1);k}^{(j)}\right] - \mathbb{E}\left[X_{L(n);k}^{(j)}\right] \right\} \\ &\quad - \frac{\theta}{1-\theta} \left(\frac{k+1}{k-1}\right)^{n-1} \left\{ \mathbb{E}\left[X_{L(n-1);k-1}^{(j)}\right] - \mathbb{E}\left[X_{L(n);k-1}^{(j)}\right] \right\} + \mathbb{E}\left[X_{L(n-1);k+1}^{(j)}\right] \end{aligned}$$

Remark 2. Setting $k = 1$ in (6) and (13), deduce the recurrence relations for marginal moment generating functions of lower record values from the complementary exponential-geometric distribution.

Relations for Joint Moment Generating Function

Next, establish the explicit and recurrence relations for joint moment generating function of k -th lower record values from complementary exponential-geometric distribution.

Theorem 4. For a positive integer $k > 1$, and for $n \geq 1$,

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$$\begin{aligned} M_{X_{L(m,n)}}(t_1, t_2) &= k^n \sum_{p=0}^{\infty} \sum_{l=0}^{\infty} \sum_{q=0}^{\infty} \sum_{r=0}^{\infty} \frac{\Gamma(\beta t_1 + 1) \Gamma(\beta t_2 + 1) (p)_l}{l! p! \theta^{p+q} \Gamma(\beta t_1 + 1 - q) \Gamma(\beta t_2 + 1 - p)} \\ &\quad \times \frac{(q)_r}{r! q! (p+l+k)^{n-m} (p+l+q+r+k)^m} \end{aligned} \quad (14)$$

Proof. From (4),

$$\begin{aligned} M_{X_{L(m,n)}}(t_1, t_2) &= \frac{k^n}{(m-1)!(n-m-1)!} \int_0^{\infty} \int_0^x e^{t_1 x + t_2 y} [-\ln(F(x))]^{m-1} \frac{f(x)}{F(x)} \\ &\quad \times [-\ln(F(y)) + \ln(F(x))]^{n-m-1} [F(y)]^{k-1} f(y) dy dx \\ &= \frac{k^n}{(m-1)!(n-m-1)!} \int_0^{\infty} e^{t_1 x} [-\ln(F(x))]^{m-1} \frac{f(x)}{F(x)} \\ &\quad \times \left(\int_0^{\infty} \left[1 + \frac{F(x) e^{-w}}{\theta(1-F(x) e^{-w})} \right]^{\beta t_2} w^{n-m-1} [F(x) e^{-w}]^k dw \right) dx \\ &= \frac{k^n}{(m-1)!} \sum_{p=0}^{\infty} \sum_{l=0}^{\infty} \frac{\Gamma(\beta t_2 + 1) (p)_l}{l! p! \theta^p \Gamma(\beta t_2 + 1 - p) (p+k+l)^{n-m}} \\ &\quad \times \int_0^{\infty} e^{t_1 x} [F(x)]^{p+k+l-1} [-\ln(F(x))]^{m-1} f(x) dx \\ &= \frac{k^n}{(m-1)!} \sum_{p=0}^{\infty} \sum_{l=0}^{\infty} \sum_{q=0}^{\infty} \sum_{r=0}^{\infty} \frac{\Gamma(\beta t_2 + 1) (p)_l}{l! p! \theta^p \Gamma(\beta t_2 + 1 - p) (p+k+l)^{n-m}} \\ &\quad \times \frac{\Gamma(\beta t_1 + 1) (q)_r}{q! r! \theta^q \Gamma(\beta t_1 + 1 - q)} \int_0^{\infty} e^{-u(k+p+l+r+k)} u^{m-1} du \end{aligned}$$

where $w = \ln F(x) - \ln F(y)$ and $u = -\ln F(x)$. The result follows by the definition of complete gamma distribution function.

Remark 3. Setting $k = 1$ in (14), deduce the explicit expression for joint moment generating function of lower record value from the complementary exponential-geometric distribution.

Theorem 5. For a positive integer $1 \leq m \leq n - 2$ and $i, j = 1, 2, \dots$,

$$\begin{aligned}
 & \mathbf{M}_{X_{L(m,n)k}}^{(i,j)}(t_1, t_2) \\
 &= \frac{1}{(1+k(1-2\theta))t_2} \left[k\theta \left(\frac{k}{k-1} \right)^{n-1} \left[\mathbf{M}_{X_{L(m,n-1)k-1}}^{(i,j)}(t_1, t_2) \right. \right. \\
 & \quad \left. \left. - t_2 \mathbf{M}_{X_{L(m,n)k-1}}^{(i,j)}(t_1, t_2) - j \mathbf{M}_{X_{L(m,n)k-1}}^{(i,j-1)}(t_1, t_2) \right] + (1-\theta)k \left(\frac{k}{k-1} \right)^{n-1} \right. \\
 & \quad \left. \times \left[t_2 \mathbf{M}_{X_{L(m,n)k+1}}^{(i,j)}(t_1, t_2) - \mathbf{M}_{X_{L(m,n-1)k+1}}^{(i,j)}(t_1, t_2) + j \mathbf{M}_{X_{L(m,n)k+1}}^{(i,j-1)}(t_1, t_2) \right] \right. \\
 & \quad \left. - (j+jk(1-2\theta)) \mathbf{M}_{X_{L(m,n)k}}^{(i,j-1)}(t_1, t_2) + (1-2\theta) \mathbf{M}_{X_{L(m,n-1)k}}^{(i,j)}(t_1, t_2) \right] \quad (15)
 \end{aligned}$$

Proof. From (4), for $1 \leq m \leq n-2$ and $i, j = 1, 2, \dots$,

$$\begin{aligned}
 \mathbf{M}_{X_{L(m,n)}}(t_1, t_2) &= \frac{k^n}{(m-1)!(n-m-1)!} \int_0^\infty \int_0^x e^{t_1x+t_2y} \left[-\ln(F(x)) \right]^{m-1} \frac{f(x)}{F(x)} \\
 & \quad \times \left[-\ln(F(y)) + \ln(F(x)) \right]^{n-m-1} [F(y)]^{k-1} \\
 & \quad \times \left[\theta + (1-2\theta)F(y) - (1-\theta)[F(y)]^2 \right] dy dx \\
 &= I_5 + I_6 + I_7 \quad (16)
 \end{aligned}$$

where

$$\begin{aligned}
 I_5 &= \frac{\theta k^n}{(m-1)!(n-m-1)!} \int_0^\infty \int_0^x e^{t_1x+t_2y} \left[-\ln(F(x)) \right]^{m-1} \frac{f(x)}{F(x)} \\
 & \quad \times \left[-\ln(F(y)) + \ln(F(x)) \right]^{n-m-1} [F(y)]^{k-1} dy dx \\
 I_6 &= \frac{(1-2\theta)k^n}{(m-1)!(n-m-1)!} \int_0^\infty \int_0^x e^{t_1x+t_2y} \left[-\ln(F(x)) \right]^{m-1} \frac{f(x)}{F(x)} \\
 & \quad \times \left[-\ln(F(y)) + \ln(F(x)) \right]^{n-m-1} [F(y)]^k dy dx \\
 I_7 &= \frac{(1-\theta)k^n}{(m-1)!(n-m-1)!} \int_0^\infty \int_0^x e^{t_1x+t_2y} \left[-\ln(F(x)) \right]^{m-1} \frac{f(x)}{F(x)} \\
 & \quad \times \left[-\ln(F(y)) + \ln(F(x)) \right]^{n-m-1} [F(y)]^{k+1} dy dx
 \end{aligned}$$

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By integrating by parts,

$$\begin{aligned}
 I_5 &= \frac{\theta(n-m-1)k^n}{t_2(m-1)!(n-m-1)!} \int_0^\infty \int_0^x e^{t_1x+t_2y} [-\ln(F(x))]^{m-1} \frac{f(x)}{F(x)} \\
 &\quad \times [-\ln(F(y)) + \ln(F(x))]^{n-m-2} [F(y)]^{k-2} f(y) dy dx \\
 &\quad - \frac{\theta(k-1)k^n}{t_2(m-1)!(n-m-1)!} \int_0^\infty \int_0^x e^{t_1x+t_2y} [-\ln(F(x))]^{m-1} \frac{f(x)}{F(x)} \\
 &\quad \times [-\ln(F(y)) + \ln(F(x))]^{n-m-1} [F(y)]^{k-2} f(y) dy dx
 \end{aligned} \tag{17}$$

$$\begin{aligned}
 I_6 &= \frac{(1-2\theta)(n-m-1)k^n}{t_2(m-1)!(n-m-1)!} \int_0^\infty \int_0^x e^{t_1x+t_2y} [-\ln(F(x))]^{m-1} \frac{f(x)}{F(x)} \\
 &\quad \times [-\ln(F(y)) + \ln(F(x))]^{n-m-2} [F(y)]^{k-1} f(y) dy dx \\
 &\quad - \frac{(1-2\theta)k^n}{t_2(m-1)!(n-m-1)!} \int_0^\infty \int_0^x e^{t_1x+t_2y} [-\ln(F(x))]^{m-1} \frac{f(x)}{F(x)} \\
 &\quad \times [-\ln(F(y)) + \ln(F(x))]^{n-m-1} [F(y)]^{k-1} f(y) dy dx
 \end{aligned} \tag{18}$$

and

$$\begin{aligned}
 I_7 &= \frac{(1-\theta)(n-m-1)k^n}{t_2(m-1)!(n-m-1)!} \int_0^\infty \int_0^x e^{t_1x+t_2y} [-\ln(F(x))]^{m-1} \frac{f(x)}{F(x)} \\
 &\quad \times [-\ln(F(y)) + \ln(F(x))]^{n-m-2} [F(y)]^k f(y) dy dx \\
 &\quad - \frac{(1-\theta)k^n}{t_2(m-1)!(n-m-1)!} \int_0^\infty \int_0^x e^{t_1x+t_2y} [-\ln(F(x))]^{m-1} \frac{f(x)}{F(x)} \\
 &\quad \times [-\ln(F(y)) + \ln(F(x))]^{n-m-1} [F(y)]^k f(y) dy dx
 \end{aligned} \tag{19}$$

Combining equations (16) and (17)-(19),

$$\begin{aligned}
 & \mathbf{M}_{X_{L(m,n);k}}(t_1, t_2) \\
 &= \frac{k\theta}{t_2} \left(\frac{k}{k-1} \right)^{n-1} \left[\mathbf{M}_{X_{L(m,n-1);k}}(t_1, t_2) - t_2 \mathbf{M}_{X_{L(m,n);k-1}}(t_1, t_2) \right] \\
 &+ \frac{(1-2\theta)}{t_2} \left[\mathbf{M}_{X_{L(m,n-1);k}}(t_1, t_2) - t_2 k \mathbf{M}_{X_{L(m,n);k}}(t_1, t_2) \right] \\
 &- \frac{(1-\theta)k}{t_2} \left(\frac{k}{k+1} \right)^{n-1} \left[\mathbf{M}_{X_{L(m,n-1);k+1}}(t_1, t_2) - t_2 \mathbf{M}_{X_{L(m,n);k+1}}(t_1, t_2) \right]
 \end{aligned} \tag{20}$$

Differentiating both the sides of (20) i times with respect to t_1 and then j times with respect to t_2 ,

$$\begin{aligned}
 \mathbf{M}_{X_{L(m,n);k}}^{(i,j)}(t_1, t_2) &= \frac{k\theta}{t_2} \left(\frac{k}{k-1} \right)^{n-1} \left[\mathbf{M}_{X_{L(m,n-1);k-1}}^{(i,j)}(t_1, t_2) - t_2 \mathbf{M}_{X_{L(m,n);k-1}}^{(i,j)}(t_1, t_2) \right] \\
 &+ \frac{(1-2\theta)}{t_2} \left[\mathbf{M}_{X_{L(m,n-1);k}}^{(i,j)}(t_1, t_2) - t_2 k \mathbf{M}_{X_{L(m,n);k}}^{(i,j)}(t_1, t_2) \right] \\
 &- \frac{(1-\theta)k}{t_2} \left(\frac{k}{k+1} \right)^{n-1} \left[\mathbf{M}_{X_{L(m,n-1);k+1}}^{(i,j)}(t_1, t_2) - t_2 \mathbf{M}_{X_{L(m,n);k+1}}^{(i,j)}(t_1, t_2) \right] \\
 &- \frac{jk\theta}{t_2} \left(\frac{k}{k-1} \right)^{n-1} \mathbf{M}_{X_{L(m,n);k-1}}^{(i,j-1)}(t_1, t_2) - k(1-2\theta)j \mathbf{M}_{X_{L(m,n);k}}^{(i,j-1)}(t_1, t_2) \\
 &+ \frac{k(1-\theta)}{t_2} \left(\frac{k}{k+1} \right)^{n-1} j \mathbf{M}_{X_{L(m,n);k+1}}^{(i,j-1)}(t_1, t_2) - j \mathbf{M}_{X_{L(m,n);k}}^{(i,j-1)}(t_1, t_2)
 \end{aligned}$$

The proof is complete.

Similarly, by using the relation in equation (*) and proceeding as above, obtain the following result:

Theorem 6. For a positive integer $1 \leq m \leq n - 2$ and $i, j = 1, 2, \dots$,

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$$\begin{aligned}
 & M_{X_{L(m,n)k}}^{(i,j)}(t_1, t_2) \\
 &= \frac{1}{t_2(1+k(1-2\theta))} \left[\frac{k^n \theta}{(k-1)^{n-1}} M_{X_{L(m,n-1)k-1}}^{(i,j)}(t_1, t_2) - \frac{k^n \theta}{(k-1)^{n-2}} \right. \\
 & \quad \times \left(t_2 M_{X_{L(m,n)k-1}}^{(i,j)}(t_1, t_2) + j M_{X_{L(m,n)k-1}}^{(i,j-1)}(t_1, t_2) \right) - j M_{X_{L(m,n)k}}^{(i,j-1)}(t_1, t_2) \quad (21) \\
 & \quad + (1-2\theta) M_{X_{L(m,n-1)k}}^{(i,j)}(t_1, t_2) - (1-\theta) k \left(\frac{k}{k+1} \right)^{n-1} M_{X_{L(m,n)k+1}}^{(i,j)}(t_1, t_2) \\
 & \quad \left. + (1-\theta) k \left(\frac{k}{k+1} \right)^{n-1} \left(t_2 M_{X_{L(m,n)k+1}}^{(i,j)}(t_1, t_2) + j M_{X_{L(m,n)k+1}}^{(i,j-1)}(t_1, t_2) \right) \right]
 \end{aligned}$$

Proof. Similar to the proof of Theorem 5.

By differentiating both sides of equations (15) and (21) with respect to t_1, t_2 and then setting $t_1 = t_2 = 0$, obtain the recurrence relations for product moments of lower k -th record values from the complementary exponential-geometric distribution.

$$\begin{aligned}
 E\left(X_{L(m,n-1)k+1}^{(i,j)}\right) &= \frac{1}{k(1-\theta)} \left(\frac{k+1}{k} \right)^{n-1} \left[k\theta \left(\frac{k}{k-1} \right)^{n-1} E\left(X_{L(m,n-1)k-1}^{(i,j)}\right) \right. \\
 & \quad + (1-2\theta)k\theta E\left(X_{L(m,n-1)k}^{(i,j)}\right) - jk\theta \left(\frac{k}{k-1} \right)^{n-1} E\left(X_{L(m,n)k-1}^{(i,j-1)}\right) \\
 & \quad - k(1-2\theta)j E\left(X_{L(m,n)k}^{(i,j-1)}\right) + kj(1-\theta) \left(\frac{k}{k+1} \right)^{n-1} E\left(X_{L(m,n)k+1}^{(i,j-1)}\right) \\
 & \quad \left. - j E\left(X_{L(m,n)k}^{(i,j-1)}\right) \right]
 \end{aligned}$$

and

$$\begin{aligned} E\left(X_{L(m,n)k}^{(i,j-1)}\right) &= \frac{1}{j(1-k(1-2\theta))} \left[\frac{\theta k^n}{(k-1)^{n-1}} E\left(X_{L(m,n-1)k-1}^{(i,j)}\right) - \frac{j\theta k^n}{(k-1)^{n-2}} E\left(X_{L(m,n)k-1}^{(i,j-1)}\right) \right. \\ &\quad + (1-2\theta) E\left(X_{L(m,n-1)k}^{(i,j)}\right) - (1-\theta) k \left(\frac{k}{k-1}\right)^{n-1} E\left(X_{L(m,n-1)k-1}^{(i,j)}\right) \\ &\quad \left. + (1-\theta) k \left(\frac{k}{k-1}\right)^{n-1} j E\left(X_{L(m,n)k+1}^{(i,j-1)}\right) \right] \end{aligned}$$

Remark 4. Setting $k = 1$ in (15) and (21), deduce the recurrence relations for joint moment generating function of lower record value from the complementary exponential-geometric distribution.

Remark 5. Equation (6) can be deduced from equation (15) by setting $j = 0$.

Numerical Results

The recurrence relations for the first and second single moments of lower record values allow us to evaluate the means and variances of all lower record values. For the computation of variances and covariances, the product moment were computed first. The diagonal elements are obtained from the recurrence relations. Next, the recurrence relation was used for the computation of the product moment of lower record values. These values were used to evaluate the covariance. When the recurrence relations to calculate the moments is used, only a few initial moments are needed to be numerically calculated. The values of means, variances and covariances were computed using R software. In Tables 1-3, we have computed the values of means for $\theta = 0.25, 0.5, 0.75$ and $\beta = 0.5(0.5)3$. From Tables 1-3, the means are decreasing with respect to n, β , and θ . In Tables 4-6, we have computed the variances and covariances for different values m and n and for different values of β and θ . Variances and covariances decrease for both β and θ values as n increases.

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Table 1. Means of lower record values for $\theta = 0.25$

n	$\beta = 0.5$	$\beta = 1.0$	$\beta = 1.5$	$\beta = 2.0$	$\beta = 2.5$	$\beta = 3.0$
1	7.498627	4.857795	3.476042	2.653920	2.120225	1.751189
2	2.278825	1.525475	1.120836	0.874487	0.711305	0.596444
3	0.779095	0.537407	0.404214	0.321322	0.265365	0.225329
4	0.295641	0.209053	0.160229	0.129257	0.108009	0.092597
5	0.122002	0.087910	0.068325	0.055711	0.046949	0.040525
6	0.053580	0.039127	0.030709	0.025225	0.021381	0.018542
7	0.024580	0.018114	0.014310	0.011813	0.010052	0.008744
8	0.011613	0.008610	0.006831	0.005658	0.004826	0.004207
9	0.005595	0.004164	0.003314	0.002750	0.002350	0.002051
10	0.002730	0.002037	0.001624	0.001350	0.001155	0.001009

Table 2. Means of lower record values for $\theta = 0.5$

n	$\beta = 0.5$	$\beta = 1.0$	$\beta = 1.5$	$\beta = 2.0$	$\beta = 2.5$	$\beta = 3.0$
1	2.293903	1.583854	1.191708	0.947316	0.782197	0.663994
2	0.787023	0.555218	0.424690	0.342004	0.285360	0.244326
3	0.298723	0.214544	0.166323	0.135340	0.113863	0.098148
4	0.123094	0.089625	0.070185	0.057553	0.048715	0.042199
5	0.053952	0.039670	0.031288	0.025796	0.021927	0.019059
6	0.024704	0.018287	0.014493	0.011993	0.010224	0.008907
7	0.011654	0.008666	0.006890	0.005715	0.004881	0.004259
8	0.005608	0.004183	0.003333	0.002769	0.002368	0.002068
9	0.002735	0.002043	0.001630	0.001356	0.001160	0.001014
10	0.001345	0.001006	0.000804	0.000669	0.000573	0.000501

Table 3. Means of lower record values for $\theta = 0.75$

n	$\beta = 0.5$	$\beta = 1.0$	$\beta = 1.5$	$\beta = 2.0$	$\beta = 2.5$	$\beta = 3.0$
1	1.284651	0.909454	0.697410	0.562694	0.470189	0.403047
2	0.461991	0.332279	0.257850	0.209968	0.176743	0.152413
3	0.182181	0.132703	0.103947	0.085254	0.072172	0.062523
4	0.077252	0.056804	0.044802	0.036937	0.031398	0.027290
5	0.034552	0.025575	0.020268	0.016771	0.014296	0.012454
6	0.016040	0.011926	0.009482	0.007865	0.006717	0.005860
7	0.007636	0.005695	0.004537	0.003769	0.003223	0.002815
8	0.003697	0.002762	0.002204	0.001833	0.001569	0.001371
9	0.001810	0.001354	0.001081	0.000900	0.000770	0.000674
10	0.000892	0.000668	0.000534	0.000445	0.000381	0.000333

Table 4. Variances and covariances of lower record values for $\theta = 0.25$

<i>m</i>	<i>n</i>	0.2	0.4	0.6	0.8	1.0	1.2
1	1	8.8540090	5.4945480	3.6235200	2.5069460	1.8030190	1.3389290
1	2	2.4780752	1.5627858	1.0459983	0.7336188	0.5343071	0.4014186
1	3	0.7809840	0.5004903	0.3398846	0.2415405	0.1780381	0.1352275
1	4	0.2755835	0.1791941	0.1232751	0.0886235	0.0660041	0.0506030
1	5	0.1072334	0.0705655	0.0490553	0.0355927	0.0267257	0.0206393
1	6	0.0450741	0.0299302	0.0209697	0.0153186	0.0115712	0.0089834
1	7	0.0200488	0.0133985	0.0094389	0.0069280	0.0052550	0.0040946
1	8	0.0092756	0.0062259	0.0044024	0.0032417	0.0024657	0.0019259
1	9	0.0044066	0.0029664	0.0021028	0.0015516	0.0011824	0.0009251
2	2	1.2504095	0.7888095	0.5281692	0.3706015	0.2700474	0.2029881
2	3	0.3723332	0.2387714	0.1622890	0.1154436	0.0851822	0.0647705
2	4	0.1244129	0.0810439	0.0558636	0.0402439	0.0300357	0.0230759
2	5	0.0461769	0.0304912	0.0212706	0.0154867	0.0116681	0.0090406
2	6	0.0186997	0.0124808	0.0087881	0.0064509	0.0048954	0.0038173
2	7	0.0080933	0.0054445	0.0038597	0.0028498	0.0021738	0.0017028
2	8	0.0036737	0.0024849	0.0017699	0.0013121	0.0010044	0.0007892
2	9	0.0017229	0.0011697	0.0008358	0.0006213	0.0004767	0.0003753
3	3	0.1857155	0.1191201	0.0809868	0.0576293	0.0425391	0.0323591
3	4	0.0588114	0.0383443	0.0264577	0.0190809	0.0142571	0.0109662
3	5	0.0207815	0.0137528	0.0096159	0.0070173	0.0052991	0.0041149
3	6	0.0169659	0.0114037	0.0080833	0.0059707	0.0045576	0.0035735
3	7	0.0071707	0.0048643	0.0034755	0.0025849	0.0019851	0.0015647
3	8	0.0032000	0.0021851	0.0015699	0.0011733	0.0009048	0.0007159
3	9	0.0014833	0.0010174	0.0007338	0.0005502	0.0004255	0.0003375
4	4	0.0637823	0.0417816	0.0289564	0.0209686	0.0157274	0.0121401
4	5	0.0218023	0.0145100	0.0101991	0.0074796	0.0056742	0.0044252
4	6	0.0082352	0.0055556	0.0039518	0.0029287	0.0022426	0.0017635
4	7	0.0033769	0.0023025	0.0016530	0.0012349	0.0009523	0.0007535
4	8	0.0014739	0.0010128	0.0007320	0.0005500	0.0004263	0.0003388
4	9	0.0006727	0.0004648	0.0003375	0.0002546	0.0001980	0.0001578
5	5	0.0114207	0.0076101	0.0053555	0.0039320	0.0029861	0.0023312
5	6	0.0041464	0.0028039	0.0019990	0.0014847	0.0011391	0.0008975
5	7	0.0016460	0.0011265	0.0008115	0.0006083	0.0004705	0.0003733
5	8	0.0007010	0.0004841	0.0003515	0.0002652	0.0002063	0.0001646
5	9	0.0003144	0.0002185	0.0001595	0.0001209	0.0000945	0.0000756
6	6	0.0021597	0.0014622	0.0010436	0.0007759	0.0005960	0.0004700
6	7	0.0008284	0.0005683	0.0004104	0.0003082	0.0002388	0.0001899
6	8	0.0003435	0.0002381	0.0001734	0.0001313	0.0001024	0.0000819
6	9	0.0001510	0.0001055	0.0000773	0.0000589	0.0000461	0.0000370
7	7	0.0004296	0.0002950	0.0002133	0.0001603	0.0001244	0.0000989
7	8	0.0001730	0.0001202	0.0000877	0.0000665	0.0000520	0.0000417
7	9	0.0000246	0.0000171	0.0000125	0.0000095	0.0000075	0.0000060
8	8	0.0000277	0.0000192	0.0000139	0.0000105	0.0000082	0.0000065
8	9	0.0000114	0.0000079	0.0000058	0.0000044	0.0000035	0.0000028

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Table 5. Variances and covariances of lower record values for $\theta = 0.5$

<i>m</i>	<i>n</i>	0.2	0.4	0.6	0.8	1.0	1.2
1	1	6.3262050	4.2259140	2.9791380	2.1898860	1.6643720	1.2998970
1	2	2.0784553	1.4052046	1.0012573	0.7430432	0.5695783	0.4482855
1	3	0.7592399	0.5187581	0.3730687	0.2791262	0.2155187	0.1707236
1	4	0.3033588	0.2090349	0.1514347	0.1140292	0.0885412	0.0704888
1	5	0.1299449	0.0901056	0.0656297	0.0496502	0.0387101	0.0309287
1	6	0.0585454	0.0407759	0.0298118	0.0226266	0.0176909	0.0141698
1	7	0.0273160	0.0190823	0.0139869	0.0106390	0.0083340	0.0066864
1	8	0.0130487	0.0091337	0.0067061	0.0051083	0.0040066	0.0032181
1	9	0.0063313	0.0044376	0.0032617	0.0024870	0.0019522	0.0015691
2	2	1.5427515	1.0494398	0.7519190	0.5608180	0.4318625	0.3413155
2	3	0.5516305	0.3795269	0.2746518	0.2066582	0.1603867	0.1276464
2	4	0.2164982	0.1503447	0.1096799	0.0831099	0.0649019	0.0519374
2	5	0.0914716	0.0639695	0.0469494	0.0357622	0.0280550	0.0225412
2	6	0.0408051	0.0286797	0.0211389	0.0161610	0.0127186	0.0102476
2	7	0.0189087	0.0133355	0.0098577	0.0075552	0.0059587	0.0048101
2	8	0.0089909	0.0063554	0.0047071	0.0036136	0.0028541	0.0023068
2	9	0.0043491	0.0030789	0.0022833	0.0017548	0.0013873	0.0011222
3	3	0.4140999	0.2865114	0.2083795	0.1574953	0.1227234	0.0980258
3	4	0.1599436	0.1117700	0.0819941	0.0624404	0.0489783	0.0393519
3	5	0.0667386	0.0469950	0.0347018	0.0265764	0.0209499	0.0169058
3	6	0.0295019	0.0208887	0.0154968	0.0119161	0.0094262	0.0076299
3	7	0.0135844	0.0096549	0.0071857	0.0055406	0.0043933	0.0035635
3	8	0.0064316	0.0045828	0.0034181	0.0026404	0.0020970	0.0017032
3	9	0.0031022	0.0022142	0.0016538	0.0012791	0.0010169	0.0008267
4	4	0.1215066	0.0853253	0.0628627	0.0480525	0.0378190	0.0304770
4	5	0.0501210	0.0354828	0.0263240	0.0202435	0.0160162	0.0129667
4	6	0.0219689	0.0156447	0.0116648	0.0090091	0.0071545	0.0058113
4	7	0.0100557	0.0071903	0.0053797	0.0041673	0.0033179	0.0027010
4	8	0.0047416	0.0033999	0.0025497	0.0019790	0.0015783	0.0012868
4	9	0.0022809	0.0016384	0.0012306	0.0009564	0.0007636	0.0006232
5	5	0.0385091	0.0273718	0.0203772	0.0157180	0.0124690	0.0101188
5	6	0.0167432	0.0119750	0.0089621	0.0069444	0.0055307	0.0045038
5	7	0.0076200	0.0054736	0.0041115	0.0031959	0.0025522	0.0020832
5	8	0.0035790	0.0025785	0.0019416	0.0015124	0.0012100	0.0009892
5	9	0.0017171	0.0012395	0.0009349	0.0007292	0.0005841	0.0004781
6	6	0.0129911	0.0093207	0.0069945	0.0054325	0.0043354	0.0035368
6	7	0.0058790	0.0042372	0.0031919	0.0024873	0.0019906	0.0016280
6	8	0.0027505	0.0019886	0.0015019	0.0011729	0.0009406	0.0007705
6	9	0.0013161	0.0009535	0.0007214	0.0005642	0.0004530	0.0003715
7	7	0.0045980	0.0033218	0.0025075	0.0019573	0.0015689	0.0012848
7	8	0.0021427	0.0015530	0.0011755	0.0009197	0.0007387	0.0006060
7	9	0.0010225	0.0007427	0.0005632	0.0004413	0.0003549	0.0002915
8	8	0.0016862	0.0012243	0.0009280	0.0007270	0.0005846	0.0004800
8	9	0.0008024	0.0005839	0.0004435	0.0003480	0.0002802	0.0002304

Table 6. Variances and covariances of lower record values for $\theta = 0.6$

<i>m</i>	<i>n</i>	0.2	0.4	0.6	0.8	1.0	1.2
1	1	5.0738580	3.4420660	2.4600120	1.8305140	1.4065540	1.1094200
1	2	1.7295697	1.1860306	0.8557216	0.6421224	0.4971114	0.3947366
1	3	0.6522793	0.4514078	0.3282963	0.2480763	0.1932415	0.1542892
1	4	0.2672931	0.1863044	0.1363294	0.1035690	0.0810548	0.0649846
1	5	0.1166466	0.0817262	0.0600696	0.0458099	0.0359716	0.0289246
1	6	0.0532419	0.0374374	0.0276012	0.0211045	0.0166100	0.0133828
1	7	0.0250609	0.0176645	0.0130501	0.0099960	0.0078792	0.0063568
1	8	0.0250609	0.0176645	0.0130501	0.0099960	0.0078792	0.0063568
1	9	0.0120416	0.0085013	0.0062891	0.0048229	0.0038054	0.0030728
2	2	1.4711976	1.0164525	0.7382938	0.5573363	0.4338054	0.3461500
2	3	0.5497578	0.3835629	0.2809775	0.2136950	0.1674271	0.1343784
2	4	0.2236305	0.1572308	0.1159433	0.0886876	0.0698357	0.0562997
2	5	0.0970613	0.0686276	0.0508504	0.0390581	0.0308668	0.0249629
2	6	0.0441328	0.0313270	0.0232894	0.0179397	0.0142126	0.0115191
2	7	0.0207192	0.0147462	0.0109873	0.0084797	0.0067291	0.0054618
2	8	0.0099381	0.0070855	0.0052872	0.0040857	0.0032458	0.0026370
2	9	0.0048350	0.0034512	0.0025778	0.0019937	0.0015850	0.0012885
3	3	0.4788187	0.3362737	0.2477615	0.1893954	0.1490620	0.1201233
3	4	0.1938386	0.1372307	0.1018093	0.0782924	0.0619422	0.0501472
3	5	0.0838327	0.0597019	0.0445150	0.0343809	0.0273035	0.0221773
3	6	0.0380233	0.0271903	0.0203445	0.0157600	0.0125482	0.0102153
3	7	0.0178209	0.0127792	0.0095842	0.0074394	0.0059335	0.0048377
3	8	0.0085384	0.0061341	0.0046077	0.0035813	0.0028596	0.0023339
3	9	0.0041510	0.0029858	0.0022451	0.0017465	0.0013956	0.0011398
4	4	0.1726281	0.1228683	0.0915759	0.0707071	0.0561396	0.0455924
4	5	0.0744937	0.0533417	0.0399609	0.0309907	0.0247002	0.0201271
4	6	0.0337354	0.0242584	0.0182380	0.0141872	0.0113373	0.0092595
4	7	0.0157948	0.0113901	0.0085839	0.0066910	0.0053564	0.0043815
4	8	0.0075624	0.0054638	0.0041242	0.0032191	0.0025800	0.0021126
4	9	0.0036749	0.0026584	0.0020087	0.0015693	0.0012587	0.0010313
5	5	0.0676009	0.0486023	0.0365367	0.0284203	0.0227109	0.0185488
5	6	0.0305842	0.0220820	0.0166596	0.0129985	0.0104147	0.0085257
5	7	0.0143102	0.0103616	0.0078361	0.0061266	0.0049175	0.0040318
5	8	0.0068487	0.0049684	0.0037634	0.0029464	0.0023676	0.0019432
5	9	0.0033271	0.0024167	0.0018325	0.0014359	0.0011548	0.0009484
6	6	0.0281612	0.0203920	0.0154228	0.0120591	0.0096800	0.0079372
6	7	0.0131705	0.0095641	0.0072508	0.0056810	0.0045683	0.0037516
6	8	0.0063013	0.0045845	0.0034812	0.0027312	0.0021988	0.0018075
6	9	0.0030606	0.0022295	0.0016947	0.0013308	0.0010722	0.0008820
7	7	0.0122564	0.0089184	0.0067730	0.0053145	0.0042790	0.0035180
7	8	0.0058625	0.0042738	0.0032508	0.0025542	0.0020589	0.0016944
7	9	0.0028470	0.0020780	0.0015822	0.0012443	0.0010038	0.0008266
8	8	0.0054964	0.0040125	0.0030556	0.0024033	0.0019389	0.0015969
8	9	0.0026687	0.0019506	0.0014869	0.0011705	0.0009451	0.0007788

Conclusion

The lower record values for complementary exponential-geometric distribution were considered. Explicit expressions were obtained, as well as recurrence relations for the marginal and joint moment generating functions of lower of record values. The recurrence relations obtained in the paper allows us to evaluate the means, variances and covariances of all lower record values for all sample sizes in a simple recursive manner.

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